

A new approach to strong embeddings

Sourav Chatterjee

UC Berkeley

Coupling of random walks with Brownian motion

- ▶ Suppose $\varepsilon_1, \varepsilon_2, \dots$ are i.i.d. random variables with $\mathbb{E}(\varepsilon_1) = 0$ and $\mathbb{E}(\varepsilon_1^2) = 1$. For each k , let

$$S_k = \sum_{i=1}^k \varepsilon_i.$$

- ▶ Goal is to construct a standard Brownian motion $(B_t)_{t \geq 0}$ on the same probability space so as to minimize the growth rate of

$$\max_{1 \leq k \leq n} |S_k - B_k|.$$

Skorokhod embeddings

- ▶ Skorokhod (1961), Strassen (1966): Start with $(B_t)_{t \geq 0}$, and construct stopping times $T_1 \leq T_2 \leq \dots$ such that
 - ▶ B_{T_1} has the same distribution as ε_1 .
 - ▶ $B_{T_1}, B_{T_2} - B_{T_1}, B_{T_3} - B_{T_2}, \dots$ are i.i.d. In other words, $(B_{T_k})_{k \geq 1}$ has the same law as $(S_k)_{k \geq 1}$.
 - ▶ $E(T_{i+1} - T_i \mid (B_t)_{t \leq T_i}) = 1$ for each i .
- ▶ Show that $T_k \simeq k$ in an average sense, to bound $\max_{1 \leq k \leq n} |B_{T_k} - B_k|$.
- ▶ Strassen (1966) proved that

$$\max_{1 \leq k \leq n} |B_{T_k} - B_k| = O((n \log \log n)^{1/4} (\log n)^{1/2}),$$

and conjectured that this is the best possible rate under $\mathbb{E}(\varepsilon_1^2) < \infty$.
Proved by Kiefer (1969).

- ▶ Is it possible to improve, assuming stronger moment conditions on the summands?

The KMT embedding theorem

Indeed, yes. If ε_1 has a finite moment generating function in a neighborhood of zero, then one can get

$$\max_{k \leq n} |S_k - B_k| = O(\log n).$$

Moreover, this is the best possible.

Theorem (Komlós-Major-Tusnády, 1975)

Let $\varepsilon_1, \varepsilon_2, \dots$ be i.i.d. random variables with $\mathbb{E} \exp \theta |\varepsilon_1| < \infty$ for some $\theta > 0$. Let $S_k := \sum_{i=1}^k \varepsilon_i$, $k = 0, 1, \dots$ be the corresponding random walk. It is possible to construct a version of the sequence $(S_k)_{k \geq 0}$ and a standard Brownian motion $(B_t)_{t \geq 0}$ on the same probability space such that for every n and every $t \geq 0$,

$$\mathbb{P}(\max_{k \leq n} |S_k - B_k| \geq C \log n + t) \leq K e^{-\lambda t},$$

where C , K , and λ do not depend on n .

Further developments

- ▶ Numerous applications in both applied and theoretical problems.
- ▶ Invaluable for understanding fine properties of simple random walk, e.g. in the works of Dembo-Peres-Rosen-Zetouni, Lawler, etc.
- ▶ There is a different KMT theorem for empirical processes, in the same paper. Of great interest to statisticians.
- ▶ No version yet for the non-i.i.d. case.
- ▶ Original paper still considered to be very hard to read. We will give a new (simpler?) proof for Bernoulli summands.

The first step

Lemma

Suppose $\varepsilon_1, \varepsilon_2, \dots, \varepsilon_n$ are i.i.d. symmetric ± 1 -valued r.v. Let $S_n = \sum_{i=1}^n \varepsilon_i$. Then it is possible to construct a version of S_n and a gaussian r.v. Z_n with mean 0 and variance n on the same probability space such that for all $t \geq 0$,

$$\mathbb{P}(|S_n - Z_n| \geq t) \leq Ke^{-\lambda t},$$

where K and λ do not depend on n .

Question: Can we have a version of this in more complex normal approximation problems?

An abstraction

Lemma

Suppose W is a random variable with $\mathbb{E}(W) = 0$ and $\mathbb{E}(W^2) < \infty$. Let T be another random variable, defined on the same probability space as W , satisfying

$$\mathbb{E}(W\varphi(W)) = \mathbb{E}(\varphi'(W)T)$$

for all Lipschitz φ . Suppose $|T|$ is almost surely bounded by a constant. Then, given any $\sigma^2 > 0$, we can construct $Z \sim N(0, \sigma^2)$ on the same probability space such that for any $\theta \in \mathbb{R}$,

$$\mathbb{E} \exp(\theta|W - Z|) \leq 2 \mathbb{E} \exp\left(\frac{2\theta^2(T - \sigma^2)^2}{\sigma^2}\right).$$

- ▶ Key idea, inspired by Stein's method: $T \simeq \sigma^2 \implies W$ is approximately $N(0, \sigma^2)$. I call T a Stein coefficient of W .
- ▶ However, classical Stein's method can only give bounds on quantities like $\sup_{f \in \mathcal{F}} |\mathbb{E}f(W) - \mathbb{E}f(Z)|$, for various classes \mathcal{F} . The above result seems to be of a fundamentally different nature.

Examples

- ▶ Suppose X is a random variable with $\mathbb{E}(X) = 0$, $\mathbb{E}(X^2) < \infty$, and density ρ . Let

$$h(x) = \frac{\int_x^\infty y\rho(y)dy}{\rho(x)}.$$

- ▶ Then, by integration-by-parts, we have

$$\mathbb{E}(X\varphi(X)) = \mathbb{E}(\varphi'(X)h(X)).$$

Thus, $h(X)$ is a Stein coefficient for X .

- ▶ Suppose X_1, \dots, X_n are i.i.d. copies of X , and let $W = n^{-1/2} \sum_{i=1}^n X_i$. Then by the above result,

$$\begin{aligned}\mathbb{E}(W\varphi(W)) &= n^{-1/2} \sum_{i=1}^n \mathbb{E}(X_i\varphi(W)) \\ &= n^{-1} \sum_{i=1}^n \mathbb{E}(h(X_i)\varphi'(W)).\end{aligned}$$

- ▶ Thus, $n^{-1} \sum h(X_i)$ is a Stein coefficient for W .

A discrete example

- ▶ Suppose $\varepsilon_1, \dots, \varepsilon_n$ are i.i.d. symmetric ± 1 -valued r.v. Let $S_n = \sum_{i=1}^n \varepsilon_i$.

- ▶ Let $Y \sim \text{Uniform}[-1, 1]$. Let $W = S_n + Y$.

- ▶ Let

$$T = n - S_n Y + \frac{1 - Y^2}{2}.$$

- ▶ It follows from a calculation involving integration-by-parts that for all φ ,

$$\mathbb{E}(W\varphi(W)) = \mathbb{E}(\varphi'(W)T),$$

that is, T is a Stein coefficient for W .

- ▶ Letting $\sigma^2 = n$, the abstract lemma tells us that it is possible to construct $Z \sim N(0, n)$ such that

$$\mathbb{E} \exp(\theta|W - Z|) \leq 2 \mathbb{E} \exp\left(\frac{2\theta^2(T - n)^2}{n}\right).$$

Since $T = n + O(\sqrt{n})$, this proves the lemma.

A general class of examples

- ▶ Suppose $\mathbf{X} = (X_1, \dots, X_n)$ be a vector of i.i.d. standard gaussian r.v.
- ▶ Let $W = f(\mathbf{X})$, where f is absolutely continuous. Suppose $\mathbb{E}(W) = 0$.
- ▶ Let $\mathbf{X}' = (X'_1, \dots, X'_n)$ be an independent copy of \mathbf{X} .
- ▶ Let

$$T = \int_0^1 \frac{1}{2\sqrt{t}} \sum_{i=1}^n \frac{\partial f}{\partial x_i}(\mathbf{X}) \frac{\partial f}{\partial x_i}(\sqrt{1-t}\mathbf{X} + \sqrt{t}\mathbf{X}') dt.$$

- ▶ Then, one can show that T is a Stein coefficient for W .
- ▶ This can be used, e.g., to prove CLTs for linear statistics of eigenvalues of random matrices. (Different talk.)

Proof of abstract lemma: Step I

- ▶ Recall setup: (W, T) is a pair of r.v. such that for all φ ,

$$\mathbb{E}(W\varphi(W)) = \mathbb{E}(\varphi'(W)T).$$

Given σ^2 , we are trying to construct $Z \sim N(0, \sigma^2)$ such that

$$\mathbb{E} \exp(\theta|W - Z|) \leq 2 \mathbb{E} \exp\left(\frac{2\theta^2(T - \sigma^2)^2}{\sigma^2}\right).$$

- ▶ Let $h(W) = \mathbb{E}(T | W)$. Then for all φ ,

$$\mathbb{E}(W\varphi(W)) = \mathbb{E}(\varphi'(W)h(W)).$$

The function h characterizes the distribution of W via this equation.

- ▶ One can show that it suffices to construct (W, Z) such that $Z \sim N(0, \sigma^2)$ and for all $\theta > 0$,

$$\mathbb{E} \exp(\theta|W - Z|) \leq 2 \mathbb{E} \exp(2\theta^2(\sqrt{h(W)} - \sigma)^2).$$

Step II: Construction of the coupling

- ▶ Fix a function $r : \mathbb{R}^2 \rightarrow \mathbb{R}$.
- ▶ For $f \in C^2(\mathbb{R}^2)$, let

$$\mathcal{L}f(x, y) := h(x) \frac{\partial^2 f}{\partial x^2} + 2r(x, y) \frac{\partial^2 f}{\partial x \partial y} + \sigma^2 \frac{\partial^2 f}{\partial y^2} - x \frac{\partial f}{\partial x} - y \frac{\partial f}{\partial y}.$$

- ▶ Suppose there exists a probability measure μ on \mathbb{R}^2 such that for all f ,

$$\int_{\mathbb{R}^2} \mathcal{L}f \, d\mu = 0. \quad (*)$$

- ▶ We will show that **every choice of r** that allows a μ satisfying **(*)** gives rise to a **coupling** of W and Z .

Step II contd.

- ▶ Recap: we have μ such that for all f

$$(*) \quad \int_{\mathbb{R}^2} \mathcal{L}f \, d\mu = 0, \quad \text{where}$$

$$\mathcal{L}f(x, y) := h(x) \frac{\partial^2 f}{\partial x^2} + 2r(x, y) \frac{\partial^2 f}{\partial x \partial y} + \sigma^2 \frac{\partial^2 f}{\partial y^2} - x \frac{\partial f}{\partial x} - y \frac{\partial f}{\partial y}.$$

- ▶ Let $(X, Y) \sim \mu$.
- ▶ Take any $\Phi \in C^2(\mathbb{R})$, and let $\varphi = \Phi'$. Putting $f(x, y) = \Phi(x)$ in $(*)$, we get

$$\mathbb{E}(h(X)\varphi'(X) - X\varphi(X)) = 0.$$

Thus, X has the same law as W .

- ▶ Similarly, putting $f(x, y) = \Phi(y)$, we get $\mathbb{E}(Y\varphi(Y)) = \sigma^2 \mathbb{E}(\varphi'(Y))$, and thus, $Y \sim N(0, \sigma^2)$.
- ▶ Note that these deductions are independent of the choice of $r(x, y)$, as long as μ exists. **Each valid choice of $r(x, y)$ gives a coupling of W and Z .**

Step III: The key lemma

Lemma

Suppose that the matrix valued function

$A(x, y) := \begin{pmatrix} h(x) & r(x, y) \\ r(x, y) & \sigma^2 \end{pmatrix}$ is bounded, positive semidefinite, and continuous everywhere. Then there exists a probability measure μ such that for all $f \in C^2(\mathbb{R}^2)$ satisfying certain mild conditions, we have

$$(*) \quad \int_{\mathbb{R}^2} \mathcal{L}f \, d\mu = 0, \quad \text{where}$$

$$\mathcal{L}f(x, y) := h(x) \frac{\partial^2 f}{\partial x^2} + 2r(x, y) \frac{\partial^2 f}{\partial x \partial y} + \sigma^2 \frac{\partial^2 f}{\partial y^2} - x \frac{\partial f}{\partial x} - y \frac{\partial f}{\partial y}.$$

Note that

- ▶ The operator \mathcal{L} is not uniformly elliptic.
- ▶ The domain of f is unbounded.
- ▶ The functions h and r need not be Lipschitz.
- ▶ Can be solved by stochastic process techniques. Will come back to give a different, self-contained proof.

Step IV: Positive definite completion

- ▶ Thus, finding $r(x, y)$ is a problem of **positive definite completion** for the incomplete matrix

$$\begin{pmatrix} h(x) & ? \\ ? & \sigma^2 \end{pmatrix}.$$

- ▶ Intuition: the more nearly singular the completion, the tighter the coupling between the two coordinates.
- ▶ The 'most singular choice' is given by the **geometric mean**

$$r(x, y) = \sigma \sqrt{h(x)}.$$

- ▶ Another choice: $r(x, y) = h(x) \wedge \sigma^2$. This is not good: It brings us back to Skorokhod embeddings. (Will not elaborate now.)

Step V: Getting bounds

- ▶ With $r(x, y) = \sigma\sqrt{h(x)}$, we have

$$\mathcal{L}f(x, y) = h(x)\frac{\partial^2 f}{\partial x^2} + 2\sigma\sqrt{h(x)}\frac{\partial^2 f}{\partial x\partial y} + \sigma^2\frac{\partial^2 f}{\partial y^2} - x\frac{\partial f}{\partial x} - y\frac{\partial f}{\partial y}.$$

- ▶ With $f(x, y) = \frac{1}{2k}(x - y)^{2k}$, we get

$$\mathcal{L}f(x, y) = (2k - 1)(x - y)^{2k-2}(\sqrt{h(x)} - \sigma)^2 - (x - y)^{2k}.$$

- ▶ So, if $\mathbb{E}(\mathcal{L}f(X, Y)) = 0$ for all f , then

$$\begin{aligned}\mathbb{E}(X - Y)^{2k} &= (2k - 1)\mathbb{E}((X - Y)^{2k-2}(\sqrt{h(X)} - \sigma)^2) \\ &\leq (2k - 1)(\mathbb{E}(X - Y)^{2k})^{\frac{k-1}{k}}(\mathbb{E}(\sqrt{h(X)} - \sigma)^{2k})^{1/k}.\end{aligned}$$

- ▶ This gives

$$\mathbb{E}(X - Y)^{2k} \leq (2k - 1)^k \mathbb{E}(\sqrt{h(X)} - \sigma)^{2k}.$$

The proof of the lemma is completed by summing over $k \geq 1$.

Proof of the key lemma

- ▶ Let A be a continuous bounded map from \mathbb{R}^2 into the set of all 2×2 positive semidefinite matrices.
- ▶ Take any probability measure μ on \mathbb{R}^2 . Suppose $\mathbf{X} \sim \mu$, and let \mathbf{Z} be a standard gaussian random vector independent of \mathbf{X} .
- ▶ For each $\varepsilon > 0$, let $T_\varepsilon \mu$ be the law of the random vector

$$(1 - \varepsilon)\mathbf{X} + \sqrt{2\varepsilon A(\mathbf{X})}\mathbf{Z},$$

where \sqrt{A} denotes the positive definite square root of A . Then T_ε is a continuous map.

- ▶ Suppose $\|A(\mathbf{x})\| \leq b$ for all $\mathbf{x} \in \mathbb{R}^2$. Let K be the set of all probability measures on \mathbb{R}^2 satisfying

$$\int \mathbf{x} d\mu(\mathbf{x}) = 0 \quad \text{and} \quad \int \exp\langle \mathbf{u}, \mathbf{x} \rangle d\mu(\mathbf{x}) \leq \exp(b\|\mathbf{u}\|^2) \quad \text{for all } \mathbf{u} \in \mathbb{R}^2.$$

Easy to verify that K is non-empty, compact, and convex.

- ▶ Main observation: For any $0 < \varepsilon < 1$, $T_\varepsilon(K) \subseteq K$.
- ▶ So, by the Schauder-Tychonoff fixed point theorem for locally convex spaces, for every $\varepsilon \in (0, 1)$, $\exists \mu_\varepsilon \in K$ such that $T_\varepsilon \mu_\varepsilon = \mu_\varepsilon$.

Proof of key lemma contd.

- ▶ Since K is compact, $\exists \mu \in K$ such that for some sequence $\varepsilon_n \downarrow 0$, $\mu_{\varepsilon_n} \rightarrow \mu$.
- ▶ Suppose $\mathbf{X}_\varepsilon \sim \mu_\varepsilon$. Using the identity $T_\varepsilon \mu_\varepsilon = \mu_\varepsilon$, we get that for any f ,

$$\mathbb{E}f((1 - \varepsilon)\mathbf{X}_\varepsilon + \sqrt{2\varepsilon A(\mathbf{X}_\varepsilon)}\mathbf{Z}) = \mathbb{E}f(\mathbf{X}_\varepsilon).$$

- ▶ Taylor expanding the LHS around \mathbf{X}_ε , dividing both sides by ε , and letting $\varepsilon \downarrow 0$ along $\{\varepsilon_n\}$, we end up with

$$(*) \quad \int_{\mathbb{R}^2} \mathcal{L}f \, d\mu = 0, \quad \text{where}$$

$$\begin{aligned} \mathcal{L}f(x, y) := & A_{11}(x, y) \frac{\partial^2 f}{\partial x^2} + 2A_{12}(x, y) \frac{\partial^2 f}{\partial x \partial y} + A_{22}(x, y) \frac{\partial^2 f}{\partial y^2} \\ & - x \frac{\partial f}{\partial x} - y \frac{\partial f}{\partial y}. \end{aligned}$$

- ▶ Thus, μ satisfies our quest and completes the proof of the lemma.

Completing the proof of the KMT theorem

We prove the following by **induction on n** .

Theorem

There exist positive universal constants C , K and λ_0 such that the following is true. Take any integer $n \geq 2$. Suppose $\varepsilon_1, \dots, \varepsilon_n$ are exchangeable ± 1 random variables. For $k = 0, 1, \dots, n$, let $S_k = \sum_{i=1}^k \varepsilon_i$ and let

$$W_k = S_k - \frac{k}{n} S_n.$$

It is possible to construct a version of W_0, \dots, W_n and a standard Brownian bridge $(\tilde{B}_t)_{0 \leq t \leq 1}$ on the same probability space such that for any $0 < \lambda < \lambda_0$,

$$\mathbb{E} \exp\left(\lambda \max_{k \leq n} |W_k - \sqrt{n} \tilde{B}_{k/n}|\right) \leq \exp(C \log n) \mathbb{E} \exp\left(\frac{K \lambda^2 S_n^2}{n}\right).$$

How to carry out the induction

The main ingredient is the following result, which is proved using the abstract lemma. (I won't go into more details.)

Theorem

Let $\varepsilon_1, \dots, \varepsilon_n$ be n arbitrary elements of $\{-1, 1\}$. Let π be a uniform random permutation of $\{1, \dots, n\}$. For each $1 \leq k \leq n$, let

$S_k = \sum_{\ell=1}^k \varepsilon_{\pi(\ell)}$, and let

$$W_k = S_k - \frac{kS_n}{n}.$$

There exist universal constants $c > 1$ and $\theta_0 > 0$ satisfying the following. Take any $n \geq 3$, any possible value of S_n , and any $n/3 \leq k \leq 2n/3$. It is possible to construct a version of W_k and a gaussian random variable Z_k with mean 0 and variance $k(n-k)/n$ on the same probability space such that for any $\theta \leq \theta_0$,

$$\mathbb{E} \exp(\theta |W_k - Z_k|) \leq \exp\left(1 + \frac{c\theta^2 S_n^2}{n}\right).$$

A few remarks about the intuition

- ▶ If $\mathbb{E}(X\varphi(X)) = \mathbb{E}(\varphi'(X)h(X))$ for all φ , then the law of X is an invariant measure for the diffusion

$$dX_t = -X_t dt + \sqrt{2h(X_t)} dB_t.$$

- ▶ Suppose we have independent diffusions

$$dX_t^i = -X_t^i dt + \sqrt{2h(X_t^i)} dB_t^i, \quad i = 1, \dots, n.$$

- ▶ Let $W_t = n^{-1/2} \sum_{i=1}^n X_t^i$. Then

$$dW_t = -W_t dt + n^{-1/2} \sum_{i=1}^n \sqrt{2h(X_t^i)} dB_t^i.$$

- ▶ If we define another process $(\tilde{B}_t)_{t \geq 0}$ by

$$d\tilde{B}_t = \frac{\sum_{i=1}^n \sqrt{2h(X_t^i)} dB_t^i}{\sqrt{\sum_{i=1}^n 2h(X_t^i)}},$$

then \tilde{B}_t is again a standard Brownian motion, and

$$dW_t = -W_t dt + \sqrt{\frac{\sum_{i=1}^n 2h(X_t^i)}{n}} d\tilde{B}_t.$$

- ▶ Since $\mathbb{E}(h(X)) = \mathbb{E}(X^2) =: \sigma^2$, therefore $n^{-1} \sum_{i=1}^n 2h(X_t^i) \simeq 2\sigma^2$ with high probability. Thus, W_t is 'approximately' an Ornstein-Uhlenbeck process.
- ▶ If we define $dZ_t = -Z_t dt + \sqrt{2}\sigma d\tilde{B}_t$, then Z_t is an actual O-U process.
- ▶ Moreover, Z_t and W_t 'come close' at infinity. Our abstract lemma gets its hands on (W_∞, Z_∞) .

An important question

Recall:

Lemma

Suppose W is a random variable with $\mathbb{E}(W) = 0$ and $\mathbb{E}(W^2) < \infty$. Let T be another random variable, defined on the same probability space as W , satisfying

$$\mathbb{E}(W\varphi(W)) = \mathbb{E}(\varphi'(W)T)$$

for all Lipschitz φ . Suppose $|T|$ is almost surely bounded by a constant. Then, given any $\sigma^2 > 0$, we can construct $Z \sim N(0, \sigma^2)$ on the same probability space such that for any $\theta \in \mathbb{R}$,

$$\mathbb{E} \exp(\theta|W - Z|) \leq 2 \mathbb{E} \exp\left(\frac{2\theta^2(T - \sigma^2)^2}{\sigma^2}\right).$$

Does there exist a multidimensional version of this lemma?

Immediate consequence: Direct proof of KMT treating the problem as a coupling of random vectors; possibly many other implications.