

Statistics 203: Introduction to Regression and Analysis of Variance

Fixed vs. Random Effects

Jonathan Taylor



Today's class

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- Two-way ANOVA
- Random vs. fixed effects
- When to use random effects?
- Example: sodium content in beer
- One-way random effects model
- Implications for model
- One-way random ANOVA table
- Inference for μ .
- Estimating σ_{μ}^2
- Example: productivity study
- Two-way random effects model
- ANOVA tables: Two-way (random)
- Mixed effects model
- Two-way mixed effects model
- ANOVA tables: Two-way (mixed)
- Confidence intervals for variances
- Satterthwaite's procedure

- Random effects.
- One-way random effects ANOVA.
- Two-way mixed & random effects ANOVA.
- Satterthwaite's procedure.



Two-way ANOVA

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- Second generalization: more than one grouping variable.
- Two-way ANOVA model: observations:
 $(Y_{ijk}), 1 \leq i \leq r, 1 \leq j \leq m, 1 \leq k \leq n_{ij}$: r groups in first grouping variable, m groups in second and n_{ij} samples in (i, j) -“cell”:

$$Y_{ijk} = \mu + \alpha_i + \beta_j + (\alpha\beta)_{ij} + \varepsilon_{ijk}, \quad \varepsilon_{ijk} \sim N(0, \sigma^2).$$

- Constraints:
 - ◆ $\sum_{i=1}^r \alpha_i = 0$
 - ◆ $\sum_{j=1}^m \beta_j = 0$
 - ◆ $\sum_{j=1}^m (\alpha\beta)_{ij} = 0, 1 \leq i \leq r$
 - ◆ $\sum_{i=1}^r (\alpha\beta)_{ij} = 0, 1 \leq j \leq m.$



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- In ANOVA examples we have seen so far, the categorical variables are well-defined categories: below average fitness, long duration, etc.
- In some designs, the categorical variable is “subject”.
- Simplest example: repeated measures, where more than one (identical) measurement is taken on the same individual.
- In this case, the “group” effect α_i is best thought of as random because we only sample a subset of the entire population of subjects.



When to use random effects?

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- A “group” effect is random if we can think of the levels we observe in that group to be samples from a larger population.
- Example: if collecting data from different medical centers, “center” might be thought of as random.
- Example: if surveying students on different campuses, “campus” may be a random effect.



Example: sodium content in beer

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- How much sodium is there in North American beer? How much does this vary by brand?
- Observations: for 6 brands of beer, researchers recorded the sodium content of 8 12 ounce bottles.
- Questions of interest: what is the “grand mean” sodium content? How much variability is there from brand to brand?
- “Individuals” in this case are brands, repeated measures are the 8 bottles.



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- Suppose we take n identical measurements from r subjects.
- $Y_{ij} \sim \mu. + \alpha_i + \varepsilon_{ij}, 1 \leq i \leq r, 1 \leq j \leq n$
- $\varepsilon_{ij} \sim N(0, \sigma^2), 1 \leq i \leq r, 1 \leq j \leq n$
- $\alpha_i \sim N(0, \sigma_{\mu}^2), 1 \leq i \leq r.$
- We might be interested in the population mean, $\mu.$: CIs, is it zero? etc.
- Alternatively, we might be interested in the variability across subjects, σ_{μ}^2 : CIs, is it zero?



Implications for model

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- In random effects model, the observations are no longer independent (even if ε 's are independent). In fact

$$\text{Cov}(Y_{ij}, Y_{i'j'}) = \sigma_{\mu}^2 \delta_{i,i'} + \sigma^2 \delta_{j,j'}.$$

- In more complicated mixed effects models, this makes MLE more complicated: not only are there parameters in the mean, but in the covariance as well.
- In ordinary least squares regression, the only parameter to estimate is σ^2 because the covariance matrix is $\sigma^2 I$.



One-way random ANOVA table

Source	SS	df	$E(MS)$
Treatments	$SSTR = \sum_{i=1}^r n (\bar{Y}_{i.} - \bar{Y}_{..})^2$	$r - 1$	$\sigma^2 + n\sigma_{\mu}^2$
Error	$SSE = \sum_{i=1}^r \sum_{j=1}^n (Y_{ij} - \bar{Y}_{i.})^2$	$(n - 1)r$	σ^2

- Only change here is the expectation of $SSTR$ which reflects randomness of α_i 's.
- ANOVA table is still useful to setup tests: the same F statistics for fixed or random will work here.
- Under $H_0 : \sigma_{\mu}^2 = 0$, it is easy to see that

$$\frac{MSTR}{MSE} \sim F_{r-1, (n-1)r}.$$

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Inference for $\mu.$

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- We know that $E(\bar{Y}_{..}) = \mu.$, and can show that

$$\text{Var}(\bar{Y}_{..}) = \frac{n\sigma_{\mu}^2 + \sigma^2}{rn}.$$

- Therefore,

$$\frac{\bar{Y}_{..} - \mu.}{\sqrt{\frac{SSTR}{(r-1)rn}}} \sim t_{r-1}$$

- Why $r - 1$ degrees of freedom? Imagine we could record an infinite number of observations for each individual, so that $\bar{Y}_{i.} \rightarrow \mu_{i.}$
- To learn anything about $\mu.$ we still only have r observations (μ_1, \dots, μ_r) .
- Sampling more within an individual cannot narrow the CI for $\mu.$



Estimating σ_{μ}^2

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■ From the ANOVA table

$$\sigma_{\mu}^2 = \frac{E(SSTR/(r-1)) - E(SSE/((n-1)r))}{n}.$$

■ Natural estimate:

$$S_{\mu}^2 = \frac{SSTR/(r-1) - SSE/((n-1)r)}{n}$$

■ Problem: this estimate can be negative! One of the difficulties in random effects model.



Example: productivity study

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- Imagine a study on the productivity of employees in a large manufacturing company.
- Company wants to get an idea of daily productivity, and how it depends on which machine an employee uses.
- Study: take m employees and r machines, having each employee work on each machine for a total of n days.
- As these employees are not *all* employees, and these machines are not *all* machines it makes sense to think of both the effects of machine and employees (and interactions) as random.



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- $Y_{ijk} \sim \mu_{..} + \alpha_i + \beta_j + (\alpha\beta)_{ij} + \varepsilon_{ijk}, 1 \leq i \leq r, 1 \leq j \leq m, 1 \leq k \leq n$
- $\varepsilon_{ijk} \sim N(0, \sigma^2), 1 \leq i \leq r, 1 \leq j \leq m, 1 \leq k \leq n$
- $\alpha_i \sim N(0, \sigma_{\alpha}^2), 1 \leq i \leq r.$
- $\beta_j \sim N(0, \sigma_{\beta}^2), 1 \leq j \leq m.$
- $(\alpha\beta)_{ij} \sim N(0, \sigma_{\alpha\beta}^2), 1 \leq j \leq m, 1 \leq i \leq r.$
- $\text{Cov}(Y_{ijk}, Y_{i'j'k'}) = \delta_{ii'}\sigma_{\alpha}^2 + \delta_{jj'}\sigma_{\beta}^2 + \delta_{ii'}\delta_{jj'}\sigma_{\alpha\beta}^2 + \delta_{ii'}\delta_{jj'}\delta_{kk'}\sigma^2.$



ANOVA tables: Two-way (random)

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SS	df	$E(SS)$
$SSA = nm \sum_{i=1}^r (\bar{Y}_{i..} - \bar{Y}...)^2$	$r - 1$	$\sigma^2 + nm\sigma_{\alpha}^2 + n\sigma_{\alpha\beta}^2$
$SSB = nr \sum_{j=1}^m (\bar{Y}_{.j.} - \bar{Y}...)^2$	$m - 1$	$\sigma^2 + nr\sigma_{\beta}^2 + n\sigma_{\alpha\beta}^2$
$SSAB = n \sum_{i=1}^r \sum_{j=1}^m (\bar{Y}_{ij.} - \bar{Y}_{i..} - \bar{Y}_{.j.} + \bar{Y}...)^2$	$(m - 1)(r - 1)$	$\sigma^2 + n\sigma_{\alpha\beta}^2$
$SSE = \sum_{i=1}^r \sum_{j=1}^m \sum_{k=1}^n (Y_{ijk} - \bar{Y}_{ij.})^2$	$(n - 1)ab$	σ^2

■ To test $H_0 : \sigma_{\alpha}^2 = 0$ use SSA and $SSAB$.

■ To test $H_0 : \sigma_{\alpha\beta}^2 = 0$ use $SSAB$ and SSE .



Mixed effects model

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- **Mixed effects model**
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- In some studies, some factors can be thought of as fixed, others random.
- For instance, we might have a study of the effect of a standard part of the brewing process on sodium levels in the beer example.
- Then, we might think of a model in which we have a fixed effect for “brewing technique” and a random effect for beer.



Two-way mixed effects model

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- $Y_{ijk} \sim \mu_{..} + \alpha_i + \beta_j + (\alpha\beta)_{ij} + \varepsilon_{ijk}, 1 \leq i \leq r, 1 \leq j \leq m, 1 \leq k \leq n$
- $\varepsilon_{ijk} \sim N(0, \sigma^2), 1 \leq i \leq r, 1 \leq j \leq m, 1 \leq k \leq n$
- $\alpha_i \sim N(0, \sigma_{\alpha}^2), 1 \leq i \leq r.$
- $\beta_j, 1 \leq j \leq m$ are constants.
- $(\alpha\beta)_{ij} \sim N(0, (m-1)\sigma_{\alpha\beta}^2/m), 1 \leq j \leq m, 1 \leq i \leq r.$
- Constraints:
 - ◆ $\sum_{j=1}^m \beta_j = 0$
 - ◆ $\sum_{i=1}^r (\alpha\beta)_{ij} = 0, 1 \leq j \leq m.$
 - ◆ $\text{Cov}((\alpha\beta)_{ij}, (\alpha\beta)_{i'j'}) = -\sigma_{\alpha\beta}^2/m$
- $\text{Cov}(Y_{ijk}, Y_{i'j'k'}) = \delta_{jj'} \left(\sigma_{\beta}^2 + \delta_{ii'} \frac{m-1}{m} \sigma_{\alpha\beta}^2 - (1 - \delta_{ii'}) \frac{1}{m} \sigma_{\alpha\beta}^2 + \delta_{ii'} \delta_{kk'} \sigma^2 \right)$



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SSB	$m - 1$	$\sigma^2 + nr \frac{\sum_{j=1}^m \beta_j^2}{m-1} + n\sigma_{\alpha\beta}^2$
$SSAB$	$(m - 1)(r - 1)$	$\sigma^2 + n\sigma_{\alpha\beta}^2$
$SSE = \sum_{i=1}^r \sum_{j=1}^m \sum_{k=1}^n (Y_{ijk} - \bar{Y}_{ij.})^2$	$(n - 1)ab$	σ^2

- To test $H_0 : \sigma_\alpha^2 = 0$ use SSA and SSE .
- To test $H_0 : \beta_1 = \dots = \beta_m = 0$ use SSB and $SSAB$.
- To test $H_0 : \sigma_{\alpha\beta}^2$ use $SSAB$ and SSE .



Confidence intervals for variances

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- Consider estimating σ_{β}^2 in the two-way random effects ANOVA. A natural estimate is

$$\widehat{\sigma}_{\beta}^2 = nr(MSB - MSAB).$$

- What about CI?
- A linear combination of χ^2 – but not χ^2 .
- To form a confidence interval for $\widehat{\sigma}_{\beta}^2$ we need to know distribution of a linear combination of MS 's, at least approximately.



Satterwaite's procedure

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- Given k independent MS 's

$$\widehat{L} \sim \sum_{i=1}^k c_i MS_i$$

- Then

$$\frac{df_T \widehat{L}}{\mathbb{E}(\widehat{L})} \text{ " } \sim \text{ " } \chi_{df_T}^2.$$

where

$$df_T = \frac{\left(\sum_{i=1}^k c_i MS_i \right)^2}{\sum_{i=1}^k c_i^2 MS_i^2 / df_i}$$

where df_i are the degrees of freedom of the i -th MS .

- $(1 - \alpha) \cdot 100\%$ CI for $\mathbb{E}(\widehat{L})$:

$$L_L = \frac{df_T \times \widehat{L}}{\chi_{df_T; 1-\alpha/2}^2}, \quad L_U = \frac{df_T \times \widehat{L}}{\chi_{df_T; \alpha/2}^2}$$